### Fundación **BBVA**

Documentos de Trabajo

2005

2

Santiago Carbó Valverde Juan Fernández de Guevara Radoselovics David Humphrey Joaquín Maudos Villarroya

# Estimating the intensity of price and non-price competition in banking:

An application to the Spanish case

### Fundación **BBVA**

### Estimating the intensity of price and non-price competition in banking: An application to the Spanish case

Santiago Carbó Valverde<sup>1</sup> Juan Fernández de Guevara Radoselovics<sup>2</sup> David Humphrey<sup>3</sup> Joaquín Maudos Villarroya<sup>2,4</sup>

UNIVERSITY OF GRANADA AND FUNCAS' INSTITUTO VALENCIANO DE INVESTIGACIONES ECONÓMICAS (IVIE)<sup>2</sup> FLORIDA STATE UNIVERSITY<sup>3</sup> UNIVERSITY OF VALENCIA<sup>4</sup>

### Abstract

We model bank oligopoly behaviour using price and non-price competition as strategic variables in an expanded conjectural variations framework. Rivals can respond to changes in both loan and deposit market prices as well as (non-price) branch market shares. The model is illustrated using data for Spain which, over 1986-2002, eliminated interest rate and branching restrictions and set off a competitive race to lock in expanded market shares. Banks use both interest rates and branches as strategic variables and both have changed over time. We illustrate the results using a regional vs. a national specification for the relevant markets.

### Resumen

En este trabajo se desarrolla un modelo de competencia bancaria oligopolista utilizando los precios y elementos distintos al precio (oficinas) como variables estratégicas en un marco de variaciones conjeturales. Los bancos rivales pueden responder a cambios en los tipos de interés de préstamos y depósitos, así como a cambios en la red de oficinas. El modelo se estima con datos del sector bancario español en el periodo 1986-2002, en el que se eliminaron las restricciones a la fijación de tipos de interés y a la expansión de oficinas. Los bancos utilizan tanto los tipos de interés como las oficinas como variables estratégicas. El modelo se estima utilizando tanto una aproximación nacional como regional a los mercados relevantes.

### Key words

non-price competition, banking, market shares.



competencia distinta de precios, banca, cuota de mercado.

La decisión de la Fundación BBVA de publicar el presente documento de trabajo no implica responsabilidad alguna sobre su contenido ni sobre la inclusión, dentro del mismo, de documentos o información complementaria facilitada por los autores.

The BBVA Foundation's decision to publish this working paper does not imply any responsibility for its content, or for the inclusion therein of any supplementary documents or information facilitated by the authors.

No se permite la reproducción total o parcial de esta publicación, incluido el diseño de la cubierta, ni su incorporación a un sistema informático, ni su transmisión por cualquier forma o medio, sea electrónico, mecánico, reprográfico, fotoquímico, óptico, de grabación u otro sin permiso previo y por escrito del titular del *copyright*.

No part of this publication, including the cover design, may be reproduced, stored in a retrieval system or transmitted in any form or by any means, electronic, mechanical, photo-copying, recording or otherwise, without the prior written permission of the copyright holder.

La serie de Documentos de trabajo, así como información sobre otras publicaciones de la Fundación BBVA pueden consultarse en: http://www.fbbva.es

#### Estimating the intensity of price and non-price competition in banking: an application to the Spanish case

EDITA © Fundación BBVA. Plaza de San Nicolás, 4. 48005 Bilbao DISEÑO DE CURIERTA

Roberto Turégano

DEPÓSITO LEGAL: M-16.564-2005 IMPRIME: Sociedad Anónima de Fotocomposición

La serie Documentos de Trabajo de la Fundación BBVA está elaborada con papel 100% reciclado, fabricado a partir de fibras celulósicas recuperadas (papel usado) y no de celulosa virgen, cumpliendo los estándares medioambientales exigidos por la actual legislación.

El proceso de producción de este papel se ha realizado conforme a las regulaciones y leyes medioambientales europeas y ha merecido los distintivos Nordic Swan y Ángel Azul.

### CONTENTS

| 1. Introduction   | 5  |
|---|----|
| 2. A model of price and non-price competition                   | 8  |
| 2.1. Basic relationships  | 8  |
| 2.2. Empirical specification                                    | 10 |
| 3. Spanish banking: competitive background and data description | 14 |
| 3.1. Spanish banking during 1986-2002                           | 14 |
| 3.2. Data and empirical approach                                | 16 |
| 4. Estimation results: price and non-price effects              | 22 |
| 4.1. Price and non-price effects on loans and deposits          | 22 |
| 4.2. The intensity of price and non-price competition           | 24 |
| 4.3. Results after all deregulation was in place                | 25 |
| 5. Summary and conclusions                                      | 29 |
| References  | 32 |
| Appendix  | 34 |

### 1. Introduction

ALMOST all empirical analyses of competition in banking in Europe and elsewhere focus on indicators of industry price competition to guide their antitrust and merger policies. In addition to long-standing efforts to divine existing and possible future price competition from measures of deposit or loan market structure, direct measures - such as the Panzar and Rosse (1987) H-statistic, loan or deposit interest margins, and Lerner indices – are increasingly relied upon as ancillary information. The H-statistic ranks current competitive behaviour on a scale from 1.0 (perfect competition) to less than or equal to 0.0 (monopoly) based upon the degree to which changes in input prices are reflected in contemporaneous changes in unit revenues. While intermediate values can signal more or less competition, there is no guideline regarding the point at which a sufficiently competitive market becomes an insufficient one. As studies by De Bandt and Davis (2000), Bikker and Haaf (2002), Carbó, Humphrey and Rodriguez (2003a) and Maudos and Pérez (2003) all find evidence of (intermediate) monopolistic behaviour for European banking markets, this information is most useful when combined with other indicators of competition.

More direct information is contained in interest margins and Lerner indices which estimate the average mark-up of price over unit cost and so indicate the current level or change in unit profitability. Corvosier and Gropp (2002) analyse the effect of concentration on margins in European banking during the 1990s and find increasing concentration is associated with less competitive pricing of loans and sight deposits but greater price competition for savings and time deposits. Similarly, Fernández de Guevara, Maudos, and Pérez (2005) estimate Lerner indexes for the 1990s and find that market power in major European countries has apparently not declined despite a series of market liberating measures. Finally, Maudos and Fernández de Guevara (2004) analyse margins and market power in major European banking sectors and illustrate the importance of including deposit and loan production costs in the margin definition. More recently, these price-based indicators of competition have been augmented with non-price measures of competitive behaviour under the assumption that banks may substitute one for the other in certain instances. For example, Pinho (2000) looks at advertising expenditures and branches as non-price strategic variables for Portugal, Kim and Vale (2001) focus on branches and their effect on loan market competition in Norway, Cesari and Chizzolini (2002) do the same for the deposit markets in Europe, while Barros (1995, 1999) uses differences in regional markets in Portugal as a strategic variable in the bank loan/deposit pricing decision.

While Kim and Vale (2001), Canhoto (2004) and Coccorese (2004) focus on rivals' responses in the loan market, and Barros (1999) and Pinho (2000) focus on similar responses in the deposit market, we try to determine their separate effects and relative intensity by looking at both markets simultaneously (rather than in isolation). Potentially, there is a four-way trade-off between changes and rivals' response in deposit and loan pricing, as well as non-price strategic efforts to alter deposit and loan market shares through de novo branch entry or acquisition. Although Kim and Vale (2001) specify that all rivals' responses occur in a national market for loans and Coccorese (2004) specifies a national market for only the largest banks in Italy, we illustrate our results using first a regional and then a national specification for the relevant deposit and loan markets. Within each market area, loan and deposit interest rates and branch network structure are determined simultaneously based on exogenous information and the likely response of rivals.

The model is illustrated using data for the Spanish banking system during 1986-2002. In anticipation of expanded competition following Spain's entry into the European Union in 1986, restrictions on bank interest rates and geographical controls on branching were removed. This permitted banks to set deposit and loan rates in response to market conditions and to compete for deposit market share and loan relationships using branches as an additional strategic variable to their pricing decisions. As a result, price and non-price behaviour is intertwined, and we provide a way to assess their relative importance as well as to determine the effect of rivals' responses by estimating conjectural variation parameters for interest rates and branches.

#### ESTIMATING THE INTENSITY OF PRICE AND NON-PRICE COMPETITION

Our model and its empirical specification are presented in Section 2, while Section 3 briefly notes key price and non-price features of the Spanish banking sector and outlines the data and empirical approach to implement our model. Empirical results are discussed in Section 4, while a summary and conclusions are presented in Section 5.

### 2. A model of price and non-price competition

#### 2.1. Basic relationships

Following Freixas and Rochet (1997), we apply a static equilibrium model with product differentiation where banks can compete with rivals in prices for deposits and loans as well as branches (our non-price variable). There are *n* banks (*i* = 1,..., *n*) and the markets for deposits (*D*) and loans (*L*) are characterised by competition in prices and product differentiation while banks are price-takers in the purchased funds or money market (*M*). With product differentiation, the demand for loans ( $l_{it}$ ) and supply of deposits ( $d_{it}$ ) at time *t* is a function of the bank's own ( $r_{it}^{l}, r_{it}^{d}$ ) and its rivals' ( $r_{iRt}^{l}, r_{iRt}^{d}$ ) interest rates on these banking outputs <sup>1</sup>, the size of its own ( $b_{it}$ ) and rivals' ( $b_{iRt}$ ) branch network, and a vector of exogenous factors which may influence the overall demand for loans and deposits ( $z_{it}^{l}, z_{it}^{d}$ ):

$$l_{it} = l(b_{it}, b_{iRt}, r_{it}^l, r_{iRt}^l, z_{it}^l)$$
(2.1)

$$d_{it} = d(b_{it}, b_{iRt}, r_{it}^d, r_{iRt}^d, z_{it}^d)$$
(2.2)

Loans made by bank *i* are expected to decrease with increases in its own interest rate  $(\partial l_{it}/\partial r_{it}^{l} < r_{it}^{l} < 0)$  and expansions of rivals' branch networks  $(\partial l_{it}/\partial b_{iRt} < 0)$ , but rise with growth in its own branch network

<sup>&</sup>lt;sup>1</sup> The demand for loans and supply of deposits for a specific bank depends on the interest rates of the (n-1) rival banks. With the aim of reducing the number of parameters to be estimated, we replace the (n-1) individual rivals' interest rates by a single condensed measure. This measure can be computed as a weighted average of the (n-1) rivals' interest rates:

 $r_{iRl} = \sum_{i \neq i}^{n-1} w_j r_{jl}$ , where *w* are the weights.

 $(\partial l_{it}/\partial b_t > 0)$  and increases in rivals' loan rates  $(\partial l_{it}/\partial r_{iRt}^l > 0)$ . Similarly, deposits at bank *i* are expected to rise with its own interest rate  $(\partial d_{it}/\partial r_{it}^d > 0)$  and growth in its branch network  $(\partial d_{it}/\partial b_{it} > 0)$ , but fall with increases in rivals' deposit rates  $(\partial d_{it}/\partial r_{iR}^d < 0)$  and branch network growth  $(\partial d_{it}/\partial b_{iRt} < 0)$ .

Bank production or operating costs  $c_{it}$  depend on the level of loan and deposit outputs and the prices of their factor (not funding) inputs  $(w_{it})$ :

$$c_{it} = c(l_{it}, d_{it}, w_{it}).$$
 (2.3)

Profits  $(\pi_{it})$  are determined from the difference between interest income and financial and operating costs:

$$\pi_{it} = r_{it}^{l} l_{it}(b_{it}, b_{iRt}, r_{it}^{l}, r_{iRt}^{l}, z_{it}^{l}) + M_{it}\overline{r_{t}} - r_{it}^{d} d_{it}(b_{it}, b_{iRt}, r_{it}^{d}, r_{iRt}^{d}, z_{it}^{d}) - -c_{it}(l_{it}, d_{it}, w_{it}) = (r_{it}^{l} - \overline{r}) l_{it}(b_{it}, b_{iRt}, r_{it}^{l}, r_{iRt}^{l}, z_{it}^{l}) + (2.4) + (\overline{r_{t}} - r_{it}^{d}) d_{it}(b_{it}, b_{iRt}, r_{it}^{d}, r_{iRt}^{d}, z_{it}^{d}) - c_{it}(l_{it}, d_{it}, w_{it})$$

where M = l - d is the net position in the money market, and  $\bar{r}_t$  is the money market rate.

To maximise profits, a bank determines the number of branches and loan and deposit interest rates from:

$$\frac{\partial \pi_{it}}{\partial b_{it}} = \left(r_{it}^{l} - \overline{r_{l}} - \frac{\partial c_{it}}{\partial l_{it}}\right) \left(\frac{\partial l_{it}}{\partial b_{it}} + \frac{\partial l_{it}}{\partial b_{iRt}}\frac{\partial b_{iRt}}{\partial b_{it}}\right) + \left(\overline{r_{t}} - r_{it}^{d} - \frac{\partial c_{it}}{\partial d_{it}}\right) \left(\frac{\partial d_{it}}{\partial b_{it}} + \frac{\partial d_{it}}{\partial b_{iRt}}\frac{\partial b_{iRt}}{\partial b_{it}}\right) = 0$$
(2.5)

$$\frac{\partial \pi_{it}}{\partial r_{it}^{l}} = l_{it} + \left(r_{it}^{l} - \overline{r_{t}} - \frac{\partial c_{it}}{\partial l_{it}}\right) \left(\frac{\partial l_{it}}{\partial r_{it}^{l}} + \frac{\partial l_{it}}{\partial r_{iRt}^{l}} \frac{\partial r_{iRt}^{l}}{\partial r_{it}^{l}}\right) = 0$$
(2.6)

$$\frac{\partial \pi_{it}}{\partial r_{it}^d} = -d_{it} + \left(\overline{r_t} - r_{it}^d - \frac{\partial c_{it}}{\partial d_{it}}\right) \left(\frac{\partial d_{it}}{\partial r_{it}^d} + \frac{\partial d_{it}}{\partial r_{iRt}^d} \frac{\partial r_{iRt}^d}{\partial r_{it}^d}\right) = 0$$
(2.7)

The terms in parentheses reflect the interest margin on loans  $(r^{l} - \bar{r})$ , deposits  $(\bar{r} - r^{d})$ , and their associated marginal operating costs  $(\partial c_{it}/\partial l_{it}, \partial c_{it}/\partial d_{it})$ . Own-price derivatives of demand for loans and depo-

sits are, respectively,  $\partial l_{il}/\partial r_{il}^{l}$  and  $\partial d_{il}/\partial r_{il}^{d}$ , while  $\partial l_{i}/\partial r_{il}^{l}$  and  $\partial d_{il}/\partial r_{iRl}^{d}$  represent rivals' price derivatives for the same two banking service *outputs*.

The terms  $\partial l_{it}/\partial b_{it}$  and  $\partial d_{it}/\partial b_{it}$  reflect the effect on loans and deposits for bank *i* due to changes in the number of its own branches, while the following three terms  $\partial b_{iRt} / \partial b_{it}$ ,  $\partial r_{iRt}^{l} / \partial r_{it}^{l}$  and  $\partial r_{iRt}^{d} / \partial r_{it}^{d}$  capture the price and non-price effect from rival banks' reactions. These last three terms include the conjectural variations (or conduct parameters) linking bank *i*'s behaviour to reactions by rivals. Conjectural variations may also be interpreted as a measure of the departure from Nash behaviour. In the case of interest rates, a zero value of these terms would imply that bank *i* completely ignores rival banks in making its decisions (Nash behaviour, where firms act taking rivals' prices as given), and a unit value means that bank *i* believes that rival banks exactly match its decisions (cartel behaviour). When  $\partial r_{iRt}/\partial r_{it} < 0$ , conduct is more competitive than Nash behaviour with prices approaching marginal costs as  $\partial r_{iRt} / \partial r_{it} \rightarrow -\infty$ . Collusive behaviour is consistent with  $\partial r_{iRt} / \partial r_{it} > 0$  suggesting that firms achieve market power through collusion.

#### 2.2. Empirical specification

In estimating the above model, the loan demand and deposit supply functions are specified as log-linear relationships:

$$\ln l_{it} = \phi_b^l \ln b_{it} + \phi_{bR}^l \ln b_{iRt} + \phi_r^l \ln r_{it}^l + \phi_{rR}^l \ln r_{iRt}^l + z_{it}^l$$
(2.8)

$$\ln d_{it} = \phi_b^d \ln b_{it} + \phi_{bR}^d \ln b_{iRt} + \phi_r^d \ln r_{it}^d + \phi_{rR}^d \ln r_{iRt}^d + z_{it}^d$$
(2.9)

where  $\phi_b^l, \phi_b^d$  ( $\phi_{bR}^l, \phi_{bR}^d$ ) are the elasticity effect from bank *i*'s own (rivals') branches while  $\phi_r^l, \phi_r^d$  ( $\phi_{rR}^l, \phi_{rR}^d$ ) are the loan and deposit elasticities from bank *i*'s own (rivals') loan and deposit interest rates.

Derivatives of the loan demand (2.8) and deposit supply (2.9) functions with respect to branches and interest rates for use in (2.5), (2.6) and (2.7) are:

$$\frac{\partial l_{it}}{\partial b_{it}} + \frac{\partial l_{it}}{\partial b_{it}} \frac{\partial b_{iRt}}{\partial b_{it}} = l_{it} \left( \frac{1}{b_{it}} \phi_b^l + \frac{1}{b_{Rt}} \phi_{bR}^l \alpha^b \right)$$
(2.10)

ESTIMATING THE INTENSITY OF PRICE AND NON-PRICE COMPETITION

$$\frac{\partial d_{it}}{\partial b_{it}} + \frac{\partial d_{it}}{\partial b_{it}} \frac{\partial b_{iRt}}{\partial b_{it}} = d_{it} \left( \frac{1}{b_{it}} \phi_b^d + \frac{1}{b_{Rt}} \phi_{bR}^d \alpha^b \right)$$
(2.11)

$$\frac{\partial l_{it}}{\partial r_{it}^{l}} + \frac{\partial l_{it}}{\partial r_{iRt}^{l}} \frac{\partial r_{iRt}^{l}}{\partial r_{it}^{l}} = l_{it} \left( \frac{\phi_{r}^{l}}{r_{it}^{l}} + \frac{\phi_{rR}^{l}}{r_{iRt}^{l}} \alpha^{l} \right) = \frac{l_{it}}{r_{it}^{l}} \left( \phi_{r}^{l} + \phi_{rR}^{l} \frac{r_{it}^{l}}{r_{iRt}^{l}} \alpha^{l} \right)$$
(2.12)

$$\frac{\partial d_{it}}{\partial r_{it}^d} + \frac{\partial d_{it}}{\partial r_{iRt}^d} \frac{\partial r_{iRt}^d}{\partial r_{it}^d} = d_{it} \left( \frac{\phi_r^d}{r_{it}^d} + \frac{\phi_{rR}^d}{r_{iRt}^d} \alpha^d \right) = \frac{d_{it}}{r_{it}^d} \left( \phi_r^d + \phi_{rR}^d \frac{r_{it}^d}{r_{iRt}^d} \alpha^d \right)$$
(2.13)

where  $\alpha^{b} = \partial b_{iRt} / \partial b_{it}$ ,  $\alpha^{l} = \partial r_{iRt}^{l} / \partial r_{it}^{l}$ ,  $\alpha^{d} = \partial r_{iRt}^{d} / \partial r_{it}^{d}$  are the conjectural variations.

From a standard translog cost function:

$$\ln c_{it} = \gamma_{0} + \sum_{h} \gamma_{h} \ln w_{hit} + \gamma_{l} \ln l_{it} + \gamma_{d} \ln d_{it} + \frac{1}{2} \sum_{h} \sum_{m} \gamma_{hm} \ln w_{hit} \ln w_{mit} + \frac{1}{2} \gamma_{ll} (\ln l_{it})^{2} + \frac{1}{2} \gamma_{dd} (\ln d_{it})^{2} + \gamma_{ld} \ln l_{it} \ln d_{it} + \sum_{h} \gamma_{hl} \ln w_{hit} \ln l_{it} + \frac{1}{2} \gamma_{hd} \ln w_{hit} \ln d_{it} + \mu_{1} Trend + \frac{1}{2} \mu_{2} Trend^{2} + \mu_{l} Trend \ln l_{it} + \frac{1}{2} \mu_{d} Trend \ln d_{it} + \sum_{h} \mu_{h} Trend \ln w_{hit}$$

$$(2.14)$$

the marginal operating costs of loans and deposits are given by:

$$mcl_{it} = \frac{c_{it}}{l_{it}} \left( \gamma_l + \gamma_{ll} \ln l_{it} + \gamma_{ld} \ln d_{it} + \sum_h \gamma_{hl} \ln w_{hit} + \mu_l Trend \right)$$
(2.15)

$$mcd_{it} = \frac{c_{it}}{d_{it}} \left( \gamma_d + \gamma_{dd} \ln d_{it} + \gamma_{ld} \ln l_{it} + \sum_h \gamma_{hd} \ln w_{hit} + \mu_d Trend \right)$$
(2.16)

Substituting (2.10) to (2.13), (2.15), and (2.16) into (2.5) to (2.7) maximises bank i's profits from its own and rivals' decisions concerning price and non-price variables:

$$\frac{\partial \pi_{it}}{\partial b_{it}} = \left(r_{it}^{l} - \overline{r_{t}} - mcl_{it}\right) l_{it} \left(\frac{1}{b_{it}} \phi^{l}_{b} + \frac{1}{b_{Rt}} \phi^{l}_{bR} \alpha^{b}\right) + \left(\overline{r_{t}} - r_{it}^{d} - mcd_{it}\right) d_{it} \left(\frac{1}{b_{it}} \phi^{d}_{b} + \frac{1}{b_{Rt}} \phi^{d}_{bR} \alpha^{b}\right) = 0$$
(2.17)

$$\frac{\partial \pi_{it}}{\partial r_{it}^{l}} = l_{it} + \left(r_{it}^{l} - \overline{r_{t}} - mcl_{it}\right) \frac{l_{it}}{r_{it}^{l}} \left(\phi_{r}^{l} + \phi_{rR}^{l} \frac{r_{it}^{l}}{r_{iRt}^{l}} \alpha^{l}\right) = 0$$
(2.18)

$$\frac{\partial \pi_{it}}{\partial r_{it}^d} = -d_{it} + \left(\overline{r_t} - r_{it}^d - mcd_{it}\right) \frac{d_{it}}{r_{it}^d} \left(\phi_r^d + \phi_{rR}^d \frac{r_{it}^d}{r_{iRt}^d} \alpha^d\right) = 0$$
(2.19)

From this, the following relationships can be derived:

$$\frac{r_{it}^{l} - \overline{r_{t}} - mcl_{it}}{\overline{r_{t}} - r_{it}^{d} - mcd_{it}} = -\frac{d_{it}}{l_{it}} \left( \frac{\phi_{b}^{d} + \frac{b_{it}}{b_{Rt}} \phi_{bR}^{d} \alpha^{b}}{\left(\phi_{b}^{l} + \frac{b_{it}}{b_{Rt}} \phi_{bR}^{l} \alpha^{b}\right)} \right)$$
(2.20)

$$\left(\frac{r_{il}^{l} - \overline{r_{i}} - mcl_{il}}{r_{il}^{l}}\right) = \frac{-1}{\phi_{r}^{l} + \phi_{iR}^{l}} \frac{r_{il}^{l}}{r_{iRl}^{l}} \alpha^{l}$$
(2.21)

$$\left(\frac{\overline{r_t} - r_{it}^d - mcd_{it}}{r_{it}^d}\right) = \frac{1}{\phi_r^d + \phi_{rR}^d \frac{r_{it}^d}{r_{iRt}^d} \alpha^d}$$
(2.22)

The terms  $\left(\frac{r_{u}^{l} - \overline{r_{i}} - mcl_{u}}{r_{u}^{l}}\right)$  and  $\left(\frac{\overline{r_{i}} - r_{u}^{d} - mcl_{u}}{r_{u}^{d}}\right)$  are expressions of the Lerner index for loans and deposits, respectively, and indicate the relative mark-up of price over marginal cost <sup>2</sup>.

<sup>2</sup> See Freixas and Rochet (1997).

#### ESTIMATING THE INTENSITY OF PRICE AND NON-PRICE COMPETITION

In empirical implementation, the cost function (2.14) is first estimated to determine the marginal operating costs which are then used in jointly estimating the first order conditions for the number of branches (2.20) and loan (2.21) and deposit (2.22) interest rates with the loan demand (2.8) and deposit supply (2.9) functions <sup>3</sup>.

<sup>&</sup>lt;sup>3</sup> The parameters of interest are:  $\alpha^{b}$  and  $\alpha^{l}$ ,  $\alpha^{d}$  (conjectural variations in branches and interest rates),  $\phi^{l}_{b}$ ,  $\phi^{d}_{b}$ ,  $\phi^{d}_{b}$  (elasticities of loan and deposit demand w.r.t. own branches),  $\phi^{l}_{bR}$ ,  $\phi^{d}_{bR}$  (elasticities w.r.t rivals' branches), and  $\phi^{l}_{r}$ ,  $\phi^{d}_{r}$  (own-price elasticities) and  $\phi^{l}_{cR}$ ,  $\phi^{d}_{cR}$  (rivals' price elasticities). Exogenous influences ( $z_{a}$ ) specified in the demand for loans (2.8) and supply of deposits (2.9) include the size of the market for loans and deposits. For each bank this variable is constructed as a weighted average of the market size of the provinces where the bank has branches, using as weights the relative importance of each province in terms of that bank's branches. Our results do not change if a linear time trend is added to (2.8) and (2.9) to account for economic expansion over the period.

3. Spanish banking: competitive background and data description

### 3.1. Spanish banking during 1986-2002

The most important actions to deregulate Spanish banking were taken in the 1980s. Controls on domestic interest rates and restrictions on branching and foreign bank entry were effectively removed over 1986-1989 in preparation for European integration and regulatory harmonisation. Although commercial banks already had the power to branch outside their regions, savings banks did not. After branching restrictions were lifted, savings banks rushed to enter new markets by opening new branches and merging with and acquiring other institutions inside and outside of their regions. This completely altered the domestic competitive environment.

The variation in aggregate loan and deposit interest rates for Spain over 1986-2002 is shown in Figure 3.1. During this period, both commercial and savings banks adopted aggressive pricing strategies seeking to increase their market share of deposit accounts. As the initial emphasis was on attracting deposits during the late 1980s, both deposit and loan rates were relatively high, only to fall during the 1990s as pressure to expand loans – along with reduced inflation – resulted in significant reductions in loan and deposit rates and bank interest margins. The fall in interest rates led depositors to expand into mutual funds and other off-balance-sheet savings instruments which, unlike in the U.S., are almost exclusively offered by banking firms. Even so, deposits grew by over 160% in real terms during this period.

Figure 3.1 also shows the overall change in the number of branch offices, which rose by 25% over 1986-2002. However, savings and commercial banks adopted different non-price strategies since branches at



FIGURE 3.1: Interest rates and branches in the Spanish banking sector

Source: Bank of Spain.

savings banks rose by 84% while branches fell by 15% at commercial banks. Fuentelsaz and Gómez (2001) note that savings banks initially adopted a defensive strategy prior to the lifting of branch restrictions by first expanding the number of branches in their own territory and then doing the same outside their regional area.

Even though the evolution of interest rates between savings and commercial banks was quite similar, savings banks increased the share of deposits in their total funding by 28% over 1986-2002 (from 43% in 1986 to 54% in 2002) while commercial banks reduced their share by 39% (falling from 53% to 32%)<sup>4</sup>. Commercial banks also experienced a reduction in their share of the loan market so that by 2002 savings and commercial banks had almost equal shares. As savings and commercial bank interest rates were similar over the period, the gains made by savings banks in the deposit and loan markets are likely to be primarily due to non-price (branch) competition <sup>5</sup>.

<sup>&</sup>lt;sup>4</sup> Other funding sources account for the fact that the deposit portion of the funding shares do not add up to 100%.

<sup>&</sup>lt;sup>5</sup> Recent regulatory initiatives, such as the Financial Services Action Plan (FSAP) of the European Commission, have the potential to affect bank price and non-price competition.

#### **3.2.** Data and empirical approach

Our unbalanced panel data covers more than 90% of bank assets in Spain and contains 2,194 observations over a 17-year period. Banks with missing data needed for estimating our model and some where data errors seemed fairly likely were excluded from the sample <sup>6</sup>. Data are from the reported balance sheets and profit and loss accounts of commercial and savings banks published by the AEB (*Asociacion Española de Banca*) and the CECA (*Confederacion Española de Cajas de Ahorros*)<sup>7</sup>.

As actual bank interest rates are not reported, yearly averages of loan (deposit) interest rates for each bank were estimated from ratios of loan revenues (deposit expenses) including fee income (expenses) to outstanding loan (deposit) values. This gives an average (not marginal) interest rate but, as our model is based more on the evolution of these prices than on their absolute level, this difference should not have a significant impact on our findings <sup>8</sup>.

Marginal operating costs are calculated from estimating a translog cost function (2.14). In (2.14)  $c_i$  is each bank's operating costs while the *input* prices (*w*) are  $w_1$  = price of labor (personnel costs / number of employees) and  $w_2$ = price of capital (operating costs except personnel costs / fixed assets). A time dummy variable is specified to capture the effect of technical change. Symmetry and linear homogeneity in input prices

<sup>7</sup> Data on deposits, loans, and branches are collected from the balance sheet of each bank. Information on financial and operating (personnel and other operating) expenses are collected from the profit and loss account of each bank.

<sup>8</sup> Some support here is seen from the fact that when we compute the *aggregate* ratio of bank loan revenues plus fee income to the value of loans outstanding, the evolution of this series over time closely approximates that of the market interest rate cited by the Bank of Spain. However, if fee income is excluded from this aggregate ratio, the correspondence weakens. The money market interest rate is assumed to be equal to the one-year interbank interest rate (source: Bank of Spain).

FSAP seeks to promote greater integration of wholesale and retail financial activities in Europe, and this can affect deposit pricing since wholesale (purchased) funds are funding substitutes for deposits. Also, FSAP seeks to encourage the development and use of new technologies in delivering financial services, which can reduce the competitive benefits of having a physical (branch) presence in the competition for market shares.

 $<sup>^6</sup>$  Banks with missing data plus those with input prices and/or computed loan and deposit interest rates that were outside the interval of +/- 2.5 times the relevant standard deviation were dropped from the sample. These problems affected 36 banks. Banks with extreme values likely reflect errors in the reported data and typically were associated with small foreign banks.

restrictions are imposed. Individual fixed effects have been introduced to capture the effect of other variables specific to each bank.

Data on rivals' interest rates and branches are computed in two ways. As in Kim and Vale (2001) and Coccorese (2004), we assume that rivals' responses occur in a national market framework so rivals' interest rates are computed from the weighted average of the (n-1) rivals' interest rates. Similarly, rivals' branch network response to changes by bank *i* is determined by the sum of all bank branches in the country, excluding those of bank *i*.

However, except for some very large corporate loans and money market institutions, the intensity of competition (and consequent rivals' response) may be stronger and better locally identified within regional markets. If bank *i* is in region *p*, then the number of rival bank branches would be the total number of branches in region *p* minus the number of branches bank *i* has in region *p*. This better represents the actual rivals of any bank *i*, whether bank *i* has branches nationwide or is only located in the region being considered. Specifically, if bank *i* has branches nationwide, only those branches in region *p* would be considered in this calculation. The calculation procedure used is shown in more detail in the Appendix (which also contains a table of the mean values of our data by year). The same logic applies to determining rivals' loan and deposit interest rates. That is, bank *i*'s rivals' loan and deposit interest rates in region *p* will be a weighted average of the interest rates of only those rival banks with branches in the same region <sup>9</sup>.

Rival banks can be identified in each of 52 provinces using data on the regional distribution of branch offices provided by AEB and CECA <sup>10</sup>. As Figure 3.2 shows, in the last year of our sample (2002), 16% of banks

<sup>&</sup>lt;sup>9</sup> As noted in the text, the exception would be for very large corporate loans where the market may be considered more national than regional in scope. Unfortunately, data are not available by loan size in Spain (nor distinguished between corporate and consumer categories) so data on the value of large corporate loans are not available. Deposit competition is clearly local in nature rather than nationwide.

<sup>&</sup>lt;sup>10</sup> We are assuming that a bank's business is distributed proportionally to its branches across the different provinces. As Carbó and Rodriguez (2004) show, this is quite reasonable since a high percentage of commercial and savings banks concentrate over 90% of their business only in one region. Other studies have used a distribution of branch offices over the 52 provinces in Spain to compute indices of concentration, market size, etc.: Fuentelsaz (1996), Maudos (1998 and 2001), Carbó, Humphrey and Rodriguez (2003b), and Carbó, López and Rodriguez (2003).



FIGURE 3.2: Number of provinces in which each bank has branches

Source: AEB and CECA.

have branches in more than half of the 52 provinces <sup>11</sup>. Only the four largest banking entities (three commercial banks, BBVA, BSCH and Banesto, and one savings bank, "la Caixa") have a presence in all provinces. At the other extreme, 34 commercial and saving banks have branches in only one province. With such differences in branch distribution, it is important to use a regional (provincial) approach to the measurement of rivals' branch networks and interest rates.

As shown earlier in Figure 3.1, the evolution of loan interest rates over time follows a downward pattern similar to the money market rate but falling even further. The same pattern applies to deposits, with the net result that the spread between money market purchased funds and deposits is quite small toward the end of our period. While the estimated marginal operating costs of loans in Figure 3.3 has also fallen over time – dropping by two-thirds – the marginal operating cost of deposits is rising. As a result, the ratio of loan to deposit marginal cost falls dra-

 $^{\rm n}$  Figure 3.2 uses information on all Spanish commercial and savings banks, which have decreased from 214 in 1986 to 140 in 2002.

matically from 5.6 in 1986 to 0.3 in 2002. Several things could explain this evolution of marginal costs. On the loan side, improvements in the evaluation of loan risk (credit scoring) can lower loan operating expenses while mortgage loans – which are cheaper to initiate and service – make up a larger share of loan portfolios (rising from 21% of all loans in 1986 to 55% in 2002). On the deposit side, the (smaller) rise in deposit marginal operating costs is associated with the shift of non-bank deposits on the balance sheet (which decreased from 50% in 1986 to 37% in 2002) into mutual funds. In addition, there was an 8 percentage point increase in the relative share of sight deposits, which have higher payment processing expenses than time or savings deposits. Finally, as branching restrictions were dropped for savings banks, they likely over-expanded their branch networks to gain market share (which added to operating expenses).





Source: Own elaboration from AEB, CECA and Bank of Spain.





Source: Own elaboration from AEB, CECA and Bank of Spain.

#### FIGURE 3.5: Relative margins



Source: Own elaboration from AEB, CECA and Bank of Spain.

The evolution of loan  $(r^{l} - \bar{r} - mcl)$  and deposit  $(\bar{r} - r^{d} - mcd)$  absolute margins and (Lerner index) relative margins are shown in Figures 3.4 and 3.5. In both cases, loan margins rose over the period (as loan marginal cost fell) while deposit margins fell (as deposit marginal cost rose) <sup>12</sup>. This suggests that market power may have increased in loan markets while falling in deposit markets <sup>13</sup>. Over 1997-2002, margins were negative in the deposit market, suggesting a loss leader pricing strategy. Although deposits were not a profitable product in themselves, they allowed banks to capture/maintain customers and, via this "tying arrangement", permitted the exercise of market power in the loan market.

<sup>&</sup>lt;sup>12</sup> Decomposing the change in the loan absolute margin into a change in the interest spread  $(r_i - \bar{r})$  plus a change in marginal costs over 1986 to 2002, the interest spread fell by 2.97 p.p., the marginal cost fell by 2.45 p.p., and overall the absolute margin fell by 0.52 p.p. In the case of deposits, the interest spread  $(\bar{r} - r_a)$  fell by 4.05 p.p., marginal costs increased by 0.59 p.p. and the absolute margin fell by 4.64 p.p. Consequently, in both cases the evolution of absolute margins is due more to changes in interest spreads than changes in marginal costs.

<sup>&</sup>lt;sup>13</sup> This market power result is similar to that found by Oroz and Salas (2003). These authors calculate relative margins using aggregate information on interest rates on new operations (marginal interest rates) but do not take into account marginal operating costs as we do here.

### 4. Estimation results: price and non-price effects

### 4.1. Price and non-price effects on loans and deposits

Our system of five simultaneous equations (2.8), (2.9), (2.20), (2.21) and (2.22) is estimated applying three-stage least squares using the two-step procedure noted in Section 2.2. Since some regressors are endogenous (loan and deposit interest rates and branches), we instrument them using one-period lags of the variables <sup>14</sup>.

Table 4.1 presents the results of the model using first a regional definition of loan and deposit markets and then a national market definition. For both of these markets all estimated parameters have the expected signs and are statistically significant <sup>15</sup>.

Within a regional market framework, the own price elasticity suggests that a 1% reduction in a bank's loan interest rate expands its loan volume by 1.46%, while a 1% rise in its deposit rate only expands deposit volume by 0.23%. Almost identical elasticities are found here within a national market framework.

The effect on a bank's loans and deposits from changes in rivals' interest rates mirrors that just noted for changes in a bank's own interest rate (although of course in the opposite direction). That is, a 1% rise in

<sup>&</sup>lt;sup>14</sup> An iterative non-linear program using the Gauss-Newton algorithm in TSP 4.5 is applied. Starting values were obtained from single equation estimates of (2.9) and (2.10) before estimating all the equations jointly. Parameter standard errors are robust to heteroskedasticity (Robust-White). Although the sample covers 1986-2002, one year's cross-section is lost from the panel data set by using one lagged values of the variables as instruments (resulting in 1,688 observations).

<sup>&</sup>lt;sup>15</sup> The only exception is for the branch conjectural variation parameter which is significant at the 89% level of confidence.

|  | Rivals' variables | at regional level | Rivals' variables at national level |       |  |  |
|--|-------------------|-------------------|-------------------------------------|-------|--|--|
|  | Estimate          | s.e.              | Estimate                            | s.e.  |  |  |
| Constant (loan demand equation)                              | 5.164**           | 0.226             | 5.120**                             | 0.232 |  |  |
| Elasticity of loans w.r.t. own branches $\phi_b^l$           | 0.728**           | 0.017             | 0.734**                             | 0.016 |  |  |
| Elasticity of loans w.r.t. rival branches $\phi_{bR}^{l}$    | 0.227**           | 0.035             | 0.252**                             | 0.036 |  |  |
| Elasticity of own loan interest rate $\phi_r^l$              | -1.457**          | 0.100             | -1.487**                            | 0.097 |  |  |
| Elasticity of rival loan interest rate $\phi_{R}^{l}$        | 1.117**           | 0.119             | 1.271**                             | 0.125 |  |  |
| Loan market size   | 0.104**           | 0.023             | 0.113**                             | 0.024 |  |  |
| Constant (deposit supply equation)                           | 1.730**           | 0.358             | 1.713**                             | 0.342 |  |  |
| Elasticity of deposits w.r.t. own branches $\phi_b^d$        | 0.749**           | 0.017             | 0.751**                             | 0.016 |  |  |
| Elasticity of deposits w.r.t. rival branches $\phi_{bR}^{d}$ | -0.389**          | 0.043             | -0.382**                            | 0.043 |  |  |
| Elasticity of own deposit interest rate $\phi_r^d$           | 0.230*            | 0.098             | 0.233**                             | 0.079 |  |  |
| Elasticity of rival deposit interest rate $\phi_{rB}^{d}$    | -0.463**          | 0.123             | -0.467**                            | 0.111 |  |  |
| Deposit market size  | 0.650**           | 0.041             | 0.649**                             | 0.039 |  |  |
| Conjectural variations in loan interest rate $\alpha^{t}$    | 0.901**           | 0.056             | 0.770**                             | 0.046 |  |  |
| Conjectural variations in deposit interest rate $\alpha^{d}$ | 0.810**           | 0.118             | 0.860**                             | 0.128 |  |  |
| Conjectural variations in branches $\alpha^{\flat}$          | 1.390             | 0.856             | 1.648**                             | 0.624 |  |  |
| N. obs.  | 1688              |                   | 1688                                |       |  |  |

#### Table 4.1: Empirical results (1986-2002)

\*\* Parameter significant at the 99% level of confidence; \* Parameter significant at the 95% level of confidence.

Note: standard errors computed from heteroskedastic-consistent matrix (Robust -White).

rivals' loan interest rates expands a bank's loan position by 1.12% (versus a 1.46% rise with a 1% reduction in the bank's own loan rate). For deposits, a 1% reduction in rivals' deposit rates expands a bank's deposit position by 0.46% (versus a 0.23% rise with a 1% rise in the bank's own deposit rate). Apparently, for the same 1% change there is a stronger loan response from changes in a bank's own loan rate than from that of rivals (and both elasticities are elastic), while on the deposit side changes in rivals' deposit rates generate the greater response (and both elasticities are inelastic).

These results are consistent with borrowers searching more carefully among lenders for their relatively infrequent and often large loan requests, as opposed to the case of depositors, where access to a convenient location is more highly valued due to their more frequent (sometimes multiple times a week) use of deposit banking services. Consequently, we would expect our non-price strategic variable (branches) to be more important for the deposit function than for loans.

On a regional basis, however, a bank's own branch elasticity for loans is 0.73 while that for deposits is 0.75, so expanding the number of branches at a bank by 1% expands deposits and loans by essentially the same percentage amount. This means that each new branch adds new loans at basically the same rate as it adds deposits generating a "balanced" balance sheet. These elasticities are identical (after rounding) in a national market environment.

Even so, loans and deposits are differentially affected when rivals expand their branch network. The elasticity of a bank's loans to rivals' branches within regional markets is 0.23 while that for deposits is -.39, so rivals' branches seem to positively affect a bank's own loan position but reduce its deposits. As these elasticities are, again, almost identical within a national market framework, this unexpected result for loans is not due to specifying a regional versus a national market. Most likely, the "income effect" of rising economic growth in Spain during the period, injections of previously "black money" into the economy with the need to declare Peseta holdings to obtain Euros, and falling interest rates, offset the "substitution effect" where rivals' branch expansion would be expected to take away loans from existing banks rather than add to them. Thus we believe the positive elasticity of a bank's loans to rivals' branches to perhaps be the result of a relatively rapidly expanding economy rather than a static or declining one <sup>16</sup>.

### 4.2. The intensity of price and non-price competition

Conjectural variation (or conduct) parameters reflect the intensity of price and non-price competition. The intensity of price competition in loan and deposit markets is inferred from the loan rate conjectural variation of 0.90 and that for the deposit rate of 0.81. As both of these values are significantly different from zero, Nash behavior is rejected <sup>17</sup>.

<sup>&</sup>lt;sup>16</sup> Kim and Vale (2001) only modeled the loan side and assumed a national loan market for Norway. In this framework, they found that rivals' branches had a significantly *negative* effect on a bank's loans. During their 8-year period (1988-1995) total loans in Norway grew by 21% while the number of branches fell by 20% and loans per branch expanded by approximately 51%. For the same years in Spain, loans grew by 89% and branches rose by 7% giving an approximate growth in loans per branch of 82%. This difference in loan growth may be the reason why the average bank in Spain found its loans grew even as rivals expanded their branch networks.

<sup>&</sup>lt;sup>17</sup> If we estimate the system of equations separately for commercial and savings banks, results show that market power is higher in the savings banks sector. Thus, for savings banks,

Simply put, if a given bank changes its loan (deposit) interest rate in a regional market environment, it expects that rivals will respond by changing their loan (deposit) rate by 90% (81%) of the original change. Thus the matching behavior in terms of price competition is fairly strong <sup>18</sup>. In principle, if these conduct parameters were both equal to 1.0, a bank's loan or deposit price variation would be exactly matched, leading, most likely, to an expanded reliance on a strategy of non-price competition. In this regard, strong price matching behavior is evident for years when loan rates rose or fell since separately estimated conjectural variations were the same in both cases ( $\alpha_r^l = 1.22$  for years when loan rates rose and 1.21 for years when they fell). This was not the case for deposits since price matching occurred in years when deposit rates rose ( $\alpha_r^d = 1.49$ ) but did not when rates fell (-0.68).

A common non-price strategy involves the placement of branch offices, and the estimated conjectural variation here is 1.39 in a regional market framework (1.65 with a national market). When a given bank establishes a new branch it expects its regional (national) rivals to respond by increasing their branch network by 1.39 (1.64) branches <sup>19</sup>. Judging by the larger estimated response, non-price competition in Spain appears to be more intense than price competition. Although it is easy to change interest rates, non-price competition can be less costly since, with floating rates, price competition may have a greater overall effect on deposit costs and loan revenues. Perhaps this helps to explain why branches in Spain are small and very close to one another.

### 4.3. Results after all deregulation was in place

The deregulation process in Spain was completed by 1992. Specifically, interest rates and controls on fees were liberalized in 1987; bran-

the loan and deposit rate conjectural variations are 0.49 and 1.07 respectively. For commercial banks, the loan and deposit rate conjectural variations are -0.01 and 0.07, respectively, and are not statistically different from zero. Consequently, Nash behavior is not rejected in the commercial banks sector.

<sup>&</sup>lt;sup>18</sup> In a national market environment, rivals' responses are 77% for loans and 86% for deposits.

<sup>&</sup>lt;sup>19</sup> The 1.39 figure is only significantly different from zero at the 89% level of confidence while the 1.65 value is significant at the 99% level. Either estimate is similar to but lower than the one reported in Kim and Vale (2001) for Norway (2.08).

ching restrictions were fully removed in 1989; a schedule to phase out compulsory investment requirements was approved in 1989; liquidity rules were liberalized in 1990; and capital adequacy requirements were modified in 1992. To see how our elasticity and conjectural variation results may be influenced by the use of our relatively long 17-year time period over 1986-2002, the data were divided into pre-1992 and post-1992 sub-periods and the model was re-estimated. Unfortunately, data for 1986-1992 did not permit our non-linear simultaneous equation model to converge and, when the convergence criterion was weakened, the resulting estimates contained the wrong signs. This problem may be due to the fact that deregulation was not yet complete in this earlier period. Bank competition on both a price and non-price basis was basically in its initial stages while, at the same time, a wave of mergers was occurring that destabilized the competitive reactions we are trying to estimate. Fortunately, estimation for the later period after deregulation was completed (1992-2002) was successful and the results are shown in Table 4.2.

#### Table 4.2: Empirical results (1993-2002)

|  | Rivals' variables | at regional level | Rivals' variables at national level |       |  |  |
|--|-------------------|-------------------|-------------------------------------|-------|--|--|
|  | Estimate          | s.e.              | Estimate                            | s.e.  |  |  |
| Constant (loan demand equation)                              | 0.711             | 0.428             | 0.868                               | 0.444 |  |  |
| Elasticity of loans w.r.t. own branches $\phi_b^l$           | 0.757**           | 0.021             | 0.750**                             | 0.021 |  |  |
| Elasticity of loans w.r.t. rival branches $\phi_{bR}^{l}$    | -0.818**          | 0.074             | -0.854**                            | 0.071 |  |  |
| Elasticity of own loan interest rate $\phi_r^l$              | -1.112**          | 0.114             | -0.884**                            | 0.093 |  |  |
| Elasticity of rival loan interest rate $\phi_{R}^{l}$        | 1.420**           | 0.164             | 1.239**                             | 0.159 |  |  |
| Loan market size   | 1.017 **          | 0.064             | 1.035**                             | 0.062 |  |  |
| Constant (deposit supply equation)                           | -1.602**          | 0.621             | -0.912                              | 0.567 |  |  |
| Elasticity of deposits w.r.t. own branches $\phi_b^d$        | 0.782**           | 0.023             | 0.781**                             | 0.023 |  |  |
| Elasticity of deposits w.r.t. rival branches $\phi_{bR}^{d}$ | -1.014**          | 0.080             | -0.863**                            | 0.073 |  |  |
| Elasticity of own deposit interest rate $\phi_r^d$           | 0.457**           | 0.109             | 0.349**                             | 0.088 |  |  |
| Elasticity of rival deposit interest rate $\phi_{tR}^{d}$    | -0.645**          | 0.140             | -0.618**                            | 0.141 |  |  |
| Deposit market size  | 1.182**           | 0.076             | 1.048**                             | 0.069 |  |  |
| Conjectural variations in loan interest rate $\alpha^{t}$    | 1.466**           | 0.116             | 1.366**                             | 0.135 |  |  |
| Conjectural variations in deposit interest rate $\alpha^{d}$ | 0.164             | 0.139             | 0.131                               | 0.160 |  |  |
| Conjectural variations in branches $\alpha^{\flat}$          | 0.317**           | 0.041             | 2.874**                             | 0.288 |  |  |
| N. obs.  | 958               |                   | 958                                 |       |  |  |

\*\* Parameter significant at the 99% level of confidence; \* Parameter significant at the 95% level of confidence.

Note: standard errors computed from heteroskedastic-consistent matrix (Robust-White).

The basic similarity of results between Tables 1 and 2 along with our inability to achieve reasonable estimates for the pre-1992 time period suggest that bank behavior during the post-1992 period drives the estimates for the entire period. Concentrating on the differences in results, the effect on a bank's deposits from changes in either its own or rivals' deposit rates has a somewhat greater impact in the post-1992 period, which suggests less market power <sup>20</sup>. In the loan market, own-price elasticity falls indicating greater market power. However, the effect on a bank's loan position is now larger for changes in rivals' loan rates.

In the case of branch elasticities, although the positive effects on loans and deposits from a bank's own branch expansion are equal to one another in the post-1992 period (as before), the effect from rivals is to reduce both a bank's deposits (as before) and its loans. It was previously suggested that the expansion of a rival's branch network added to a bank's loan position – effectively expanding the entire market for loans. In the post-1992 period, however, the sign is reversed so branching by rivals takes away a bank's loan market share; a result more indicative of both a slower expansion of loan demand and more effective non-price competition by rivals. Turning to conjectural variations, which reflect the intensity of price and non-price competition, within a regional market framework deposit competition appears to have increased in the post-deregulation period, as the conduct parameter falls from 0.81 for the entire period to 0.16 post-deregulation. However, competition seems to have decreased for loans (conduct parameter rises from 0.91 to 1.46). In terms of branches, the conduct parameter is significantly greater than zero (0.32), which means that banks use branches as a strategic variable <sup>21</sup>. Nash behavior is still rejected for loans and branches but not for deposits. Thus while banks still exercise some form of market power or coordination between institutions in the loan market post-1992 and rely on non-price competition using their branch networks, they now seem no longer to (significantly) consider rivals' responses when setting deposit interest rates.

<sup>&</sup>lt;sup>20</sup> As expressions (2.22) and (2.23) show, the Lerner index decreases when own-price elasticities ( $\phi_r^d$  and  $\phi_r^l$ ) increase.

<sup>&</sup>lt;sup>21</sup> The same results were obtained using a national market framework except that the branch conjectural variation rose rather than fell. We have more confidence in the regional market results as this is where we believe competition is most relevant and therefore best measured.

The evolution of the Lerner index (and changes in the interest rate conjectural variation parameter) indicates an increase of market power in the loan market but a decrease in the deposit market. The Lerner index for the loan market can be written as  $1 - (\bar{r} / r_l) - (mcl/r_l)$  permitting us to determine the relative contribution of changes in interest rates versus marginal cost in the overall change in the index. From 1986 to 2002, the contribution of interest rates ( $\bar{r} / r_l$ ) decreased 16 percentage points (from .62 to .46) whereas the contribution from marginal cost ( $mcl/r_l$ ) decreased by 11 percentage points (from .15 to .04). Thus the rise in the Lerner index for the loan market, from 0.23 in 1986 to 0.50 in 2002, is due more to changes in loan interest rate behaviour than changes in marginal costs. In the deposit market, the Lerner index fell from.55 in 1986 down to -.22 in 2002 and a similar decomposition shows that this reduction is also due more to changes in interest rates than marginal costs <sup>22</sup>.

<sup>&</sup>lt;sup>22</sup> The deposit decomposition is  $(\bar{r} / r_a) - (mcd / r_a) - 1$ . The contribution of interest rates  $(\bar{r} / r_a)$  was 48 percentage points (which fell from 1.62 to 1.14) while that from marginal costs  $(mcd/r_a)$  was 29 percentage points (which rose from .07 to .36).

## 5. Summary and conclusions

WE have estimated an expanded model of bank oligopoly behaviour by incorporating price (interest rate) and non-price (branch network) competition as strategic variables in both the market for bank loans and deposits. Conjectural variations in this expanded framework suggest that rivals can respond to changes in both loan and deposit market prices as well as through branching behaviour. Using data for Spain over 17 years (1986-2002) and for a decade after banking deregulation was complete (1992-2002) to illustrate our model, we find only a few important differences from specifying a regional market framework (common in the U.S.) versus a national one (typical in European studies). The major exception occurs in estimating branch conjectural variation (where there is an important increase at national level).

A regional market framework is felt to be more relevant, and on this level we find relatively large and elastic own-price (interest rate) elasticities in an average bank's market for loans, but small and inelastic own-price elasticities for deposits. Also, increases in rivals' loan rates are seen to add significantly to a bank's own loan position, while a reduction in rivals' deposit rates expands a bank's deposit position. The latter deposit "substitution effect" is expected, of course, but the positive effect on a bank's loans when rivals raise - not lower - their loan rate was not. As the expected "substitution effect" for loans was found when the sample was shortened to the period after deregulation was completed (1992-2002), this suggests that either a positive or negative response is possible. If the credit market is expanding rapidly enough, loan demand at the average bank may also expand even in the face of rising interest rates at rival banks. Here the overall economy-driven expansion of loans offsets the price-driven substitution effect among a bank and its rivals. In either case, however, the effect from a given price change in the loan market exceeds that in the deposit market. This is consistent with borrowers searching more carefully among lenders for their relatively infrequent and often large loan requests, compared to the case of depositors, where access to a convenient location is more highly valued due to their more frequent use of deposit banking services.

In this situation, branching – our non-price strategic variable – should be more important in the competition for deposits than for loans. While changes in a bank's own branch network affect loans and deposits almost equally, the expansion of rivals' branch networks should decrease both its loans and deposits. The expected result does occur for deposits but appears to have been offset (by rising economic growth and reduced interest rates) for loans, at least when the model is estimated for the entire 1986-2002 period. The expected result for loans, however, occurs during the sub-period 1992-2002 after deregulation was complete.

The intensity of price and non-price competition captured in our conjectural variation estimates suggests that when a given bank changes its loan rate, it expects that rivals will respond by changing their loan rate by about 90% of the original change. Similarly, changes in a bank's deposit rate generate changes in rivals' deposit rates by about 80% of the original change. Thus interest rate matching behavior seems fairly strong. While strong matching behavior exists for years when loan rates rise or fall, the response for deposit rates has been asymmetrical. There is strong matching for years when deposit rates rise (mostly after branching restrictions were lifted and savings banks were competing for market share) but weak matching for years when they fall.

The closer the deposit and loan price conjectural variation parameters are to 1.0, the more a bank would tend to rely on a strategy of non-price competition. With the current level of price competition, the establishment of a new branch by a bank leads rivals to respond by increasing their branch network by 1.39 branches. For the shorter period after deregulation was completed, strong "price matching" behavior is evident for loans (with a conduct parameter of 1.46) but less so for deposits (.16) or branches (.32). Even so, with price matching for loans non-price competition through branching becomes more important in this market.

Our results support the view that non-price competition can play an important role in banking, and that in Spain price competition has decreased in the loan market but increased in the deposit market over 1986-2002. We also find that the relative intensity of price versus non-price competition has varied over time, in our case after 1992 when the country's banking sector was finally fully deregulated. Unfortunately, such changes in price and non-price competition tend to undercut attempts to generalize to the future conjectural variation results obtained with historical information. This is not unlike trying to infer market competition from changes in market structure without knowing how entry conditions may affect this result. What this suggests is that industry measures of conjectural variation are most useful if kept up to date.

### References

- BARROS, P. (1995): "Post-entry expansion in banking: the case of Portugal", *International Journal of Industrial Organization* 13, 593-611.
- (1999): "Multi market competition in banking with an example from the case of Portuguese market", *International Journal of Industrial Organization* 17, 335-352.
- BIKKER, J. A. and HAAF, K. (2002): "Competition, concentration and their relationship: an empirical analysis of the banking industry", *Journal of Banking and Finance* 26, 2191-2214.
- CANHOTO, A. (2004): "Portuguese banking: a structural model of competition in the deposits market", *Review of Financial Economics* 13, 41-63.
- CARBÓ, S., D. HUMPHREY and F. RODRIGUEZ (2003a): "Bank deregulation is better than mergers", *Journal of International Financial Markets, Institutions and Money* 13, 429-449.
- (2003b), "Deregulation, bank competition and regional growth", *Regional Studies* 37(3), 227-237.
- R. LÓPEZ and F. RODRIGUEZ (2003): "La medición de la competencia en los mercados bancarios regionales", *Revista de Economía Aplicada* 3(9), 5-33.
- and F. RODRÍGUEZ (2004): "The finance-growth nexus: a regional perspective", European Urban and Regional Studies 11(4), 357-372, forthcoming.
- CESARI, V. and B. CHIZZOLINI (2002): "Branching and competition in the European banking industry", *Applied Economics* 34, 2213-2225.
- COCCORESE, P. (2004): "Competition in markets with dominant firms: a note on the evidence from the Italian banking industry", *Journal of Banking and Finance*, forthcoming.
- CORVOISIER, S. and R. GROPP (2002): "Bank concentration and retail interest rates", *Journal of Banking and Finance* 26, 2155-2189.
- DE BANDT, O. and E. P. DAVIS (2000): "Competition, contestability and market structure in European banking sectors on the eve of EMU", *Journal of Banking and Finance* 24, 1045-1066.
- FERNÁNDEZ DE GUEVARA, J., J. MAUDOS and F. PÉREZ (2005): "Market power in European banking", *Journal of Financial Services Research*, 27 (2), 109-138.
- FREIXAS, X. and J. ROCHET (1997): Microeconomics of Banking, Cambridge, MA: MIT Press.
- FUENTELSAZ, L. (1996): "Dinámica de la Competencia entre Cajas de Ahorros Españolas", *Investigaciones Económicas* 20, 125-41.

- FUENTELSAZ, L. and J. GÓMEZ (2001): "Strategic and queue effects at entry in Spanish banking", *Journal of Economics and Management Strategy* 10(4), 529-563.
- KIM, M. and B. VALE (2001): "Non-price strategic behaviour: the case of bank branches", *International Journal of Industrial Organisation* 19, 1583-1602.
- MAUDOS, J. (1998): "Market structure and performance in Spanish banking using a direct measure of efficiency", *Applied Financial Economics* 8, 191-200.
- (2001): "Rentabilidad, estructura de mercado y eficiencia en el sector bancario español", *Revista de Economía Aplicada* 9(25), 193-207.
- and F. PÉREZ (2003): "Competencia vs. poder de mercado en la banca española", *Moneda y Crédito* 217, 139-166.
- and J. FERNÁNDEZ DE GUEVARA (2004): "Factors explaining the interest margin in the banking sectors of the European Union", *Journal of Banking and Finance* 28/9, 2259-2281.
- OROZ, M. and V. SALAS (2003): "Competencia y eficiencia de la intermediación bancaria en España: 1977-2000", *Moneda y Crédito* 217, 73-100.
- PINHO, P. (2000): "The impact of deregulation on price and non-price competition in the Portuguese deposits market", *Journal of Banking and Finance* 24, 1515-1533.

### Appendix

**R**IVALS' branch network for bank *i* in region (province)  $p(b_{iRp})$  is calculated as:

$$b_{iRp} = \sum_{i \neq j}^{n-1} b_{jp}$$
 (A.1)

When a given bank *i* has branches in different regions, rivals' branch network for bank *i* in all regions where bank *i* is located is computed as a weighted average of rivals' branch network in each region, using as weights the regional branch distribution of bank  $i^{23}$ :

/

``

$$b_{iR} = \sum_{p} b_{iRp} \left( \frac{b_{ip}}{\sum_{p} b_{ip}} \right)$$
(A.2)

In the case of loan and deposit interest rates, rivals' interest rates in each region p are calculated as <sup>24</sup>:

$$r_{iRp} = \sum_{i \neq j}^{n-1} r_{jp} \left( \frac{b_{jp}}{\sum_{i \neq j}^{n-1} b_{jp}} \right)$$
(A.3)

<sup>&</sup>lt;sup>23</sup> We compute rivals' bank network and interest rates separately for each year. To calculate rivals' interest rates, for each bank in each year, a weighting matrix with (n-1)\*p elements is computed. Over the period 1986-2002, a matrix with n\*(n-1)\*T\*p elements (almost 20 million) is computed.

<sup>&</sup>lt;sup>24</sup> Having no information to the contrary, we assume that banks set the same interest rates across their branches.

and the rivals' interest rate for bank i in all regions is computed as a weighted average of rivals' interest rates in each province:

$$r_{iR} = \sum_{p} r_{iRp} \left( \frac{b_{ip}}{\sum_{p} b_{ip}} \right)$$
(A.4)

Descriptive statistics (means) of the variables used are in Table A1.

### Table A.1: Descriptive statistics (means)

|   | 1986   | 1987   | 1988   | 1989   | 1990   | 1991   | 1992   | 1993   | 1994   | 1995   | 1996   | 1997   | 1998   | 1999   | 2000   | 2001   | 2002   |
|---|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|
| $l_{ii}$ (loans) thousand euros         | 1,348  | 1,318  | 1,537  | 1,716  | 1,840  | 2,083  | 2,520  | 2,960  | 2,760  | 2,901  | 3,115  | 3,336  | 4,092  | 4,964  | 5,521  | 4,795  | 6,265  |
| $d_{ii}$ (deposits) thousand            |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |
| euros                                   | 1,847  | 1,862  | 2,144  | 2,263  | 2,379  | 2,578  | 3,158  | 3,505  | 3,482  | 3,588  | 3,792  | 3,932  | 4,758  | 5,785  | 6,315  | 5,510  | 6,895  |
| $r_{it}^{l}$ (loan interest rate)       | 0.182  | 0.186  | 0.172  | 0.172  | 0.177  | 0.174  | 0.164  | 0.154  | 0.125  | 0.129  | 0.122  | 0.100  | 0.088  | 0.074  | 0.078  | 0.081  | 0.070  |
| $r_{ii}^{d}$ (deposit interest rate)    | 0.077  | 0.076  | 0.071  | 0.082  | 0.090  | 0.092  | 0.084  | 0.086  | 0.062  | 0.069  | 0.063  | 0.046  | 0.036  | 0.025  | 0.032  | 0.036  | 0.029  |
| Regional $r_{iRt}^{l}$ (rivals' loan    |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |
| interest rate)                          | 0.190  | 0.196  | 0.181  | 0.182  | 0.191  | 0.181  | 0.167  | 0.158  | 0.136  | 0.138  | 0.130  | 0.105  | 0.090  | 0.078  | 0.084  | 0.086  | 0.074  |
| Regional $r_{iRt}^{d}$ (rivals' deposit |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |
| interest rate)                          | 0.069  | 0.068  | 0.063  | 0.073  | 0.084  | 0.086  | 0.081  | 0.080  | 0.061  | 0.066  | 0.061  | 0.045  | 0.037  | 0.029  | 0.035  | 0.035  | 0.029  |
| National $r_{iRt}^{l}$ (rivals' loan    |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |
| interest rate)                          | 0.177  | 0.184  | 0.174  | 0.176  | 0.186  | 0.176  | 0.163  | 0.153  | 0.131  | 0.131  | 0.125  | 0.100  | 0.089  | 0.077  | 0.084  | 0.086  | 0.075  |
| National $r_{iRt}^{d}$ (rivals' deposit |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |
| interest rate)                          | 0.072  | 0.071  | 0.066  | 0.077  | 0.088  | 0.088  | 0.082  | 0.083  | 0.065  | 0.067  | 0.064  | 0.047  | 0.039  | 0.030  | 0.037  | 0.037  | 0.030  |
| $b_{ii}$ (number of branches per        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |
| bank)                                   | 154    | 150    | 166    | 166    | 167    | 181    | 218    | 197    | 195    | 202    | 214    | 218    | 259    | 302    | 314    | 282    | 332    |
| Regional $b_{rtu}$ (rivals' branch      |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |
| network)                                | 1,386  | 1,454  | 1,424  | 1,613  | 1,762  | 1,927  | 1,740  | 1,855  | 1,899  | 2,044  | 2,063  | 2,167  | 1,972  | 1,876  | 1,948  | 1,943  | 1,960  |
| National $b_{nu}$ (rivals' branch       |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |        |
| network)                                | 27,623 | 28,009 | 28,793 | 29,566 | 30,353 | 30,999 | 31,230 | 31,689 | 31,773 | 32,551 | 33,217 | 34,053 | 34,829 | 34,889 | 34,902 | 34,372 | 34,112 |
| Number of banks                         | 177    | 175    | 169    | 179    | 178    | 171    | 141    | 148    | 146    | 149    | 144    | 145    | 125    | 116    | 112    | 107    | 103    |

Source: AEB, CECA and own elaboration.

#### ABOUT THE AUTHORS\*

**SANTIAGO CARBÓ VALVERDE**. PhD in Economics, University of Wales, Bangor (UK). Currently Professor of Economics and Head of the Economics Department at the University of Granada, and a Visiting Research Fellow at the University of Wales, Bangor (UK), he is also Head of Financial System Research at the Spanish Savings Banks Foundation (Funcas). He has served as editor and on the editorial boards of various Spanish journals. He has done consulting work for the European Commission, banking organizations and the public administration, and has published over 90 monographs, book chapters and articles dealing with issues like the financegrowth link, bank efficiency, bank regulation, bank technology and financial exclusion. Website: *http://www.ugr.es/local/scarbo* 

JUAN FERNÁNDEZ DE GUEVARA RADOSELOVICS holds a B. A. in Economic and Managerial Sciences (University of Valencia, 1995). He works at the Instituto Valenciano de Investigaciones Económicas (Ivie) and at present is writing his PhD thesis on "Efficiency and competition in banking sectors: recent changes in Spain and in Europe". His research interest is bank economics.

**DAVID HUMPHREY** holds a PhD in Economics from the University of California (Berkeley). He has been Vice President and Payments System Adviser of the Federal Reserve Bank of Richmond and Assistant Director of the Division of Research and Statistics (Board of Governors of the Federal Reserve System, Washington, DC). Currently, he is Professor of Finance at Florida State University (Florida, USA). He has published over 90 articles in specialised journals. He is a member of the editorial boards of the *Journal of Monetary Economics* and the *Journal of Banking and Finance*.

JOAQUÍN MAUDOS VILLARROYA. PhD in Economics from the University of Valencia (1995), he is currently a Lecturer in Economic Analysis at the said University and Research Professor at the Instituto Valenciano de Investigaciones Economicas (Ivie). His specialty fields are banking and regional economics. He was a visiting researcher during 1995-96 at Florida State University. He has jointly published four books and over fifty articles in specialised journals, both national and international. He is a member of the editorial board of the *European Review of Economics and Finance* and principal researcher on the project (SEJ2004-00110) "Knowledge, innovation and infrastructures" of the Spanish Ministry of Science and Technology, and the project "Integration, competition and efficiency in the European financial markets" of the BBVA Foundation (website: *http://www.uv.es/maudosj)*.

\* We gratefully acknowledge the financial support of the BBVA Foundation. We also thank an anonymous referee, Jorge Sempere, Jean Dermine, Jordi Gual and other participants in the 25th SUERF Colloquium on "Competition and Profitability in European Financial Services: Strategic, Systemic and Policy Issues", 14-16 October, 2004, held in Madrid, for their very helpful comments.

### Fundación **BBVA**

### DOCUMENTOS DE TRABAJO

#### NÚMEROS PUBLICADOS

| DT 01/02 | Trampa del desempleo y educación: un análisis de las relaciones entre los efectos<br>desincentivadores de las prestaciones en el Estado del Bienestar y la educación<br>Jorge Calero Martínez y Mónica Madrigal Bajo |
|----------|--|
| DT 02/02 | Un instrumento de contratación externa: los vales o cheques.<br>Análisis teórico y evidencias empíricas<br>Ivan Planas Miret   |
| DT 03/02 | Financiación capitativa, articulación entre niveles asistenciales<br>y descentralización de las organizaciones sanitarias<br>Vicente Ortún-Rubio y Guillem López-Casasnovas  |
| DT 04/02 | La reforma del IRPF y los determinantes de la oferta laboral<br>en la familia española<br>Santiago Álvarez García y Juan Prieto Rodríguez  |
| DT 05/02 | The Use of Correspondence Analysis in the Exploration<br>of Health Survey Data<br>Michael Greenacre  |
| DT 01/03 | ¿Quiénes se beneficiaron de la reforma del IRPF de 1999?<br>José Manuel González-Páramo y José Félix Sanz Sanz   |
| DT 02/03 | <i>La imagen ciudadana de la Justicia</i><br>José Juan Toharia Cortés  |
| DT 03/03 | Para medir la calidad de la Justicia (I): Abogados<br>Juan José García de la Cruz Herrero  |
| DT 04/03 | Para medir la calidad de la Justicia (II): Procuradores<br>Juan José García de la Cruz Herrero   |
| DT 05/03 | Dilación, eficiencia y costes: ¿Cómo ayudar a que la imagen de la Justicia<br>se corresponda mejor con la realidad?<br>Santos Pastor Prieto  |
| DT 06/03 | Integración vertical y contratación externa en los servicios<br>generales de los hospitales españoles<br>Jaume Puig-Junoy y Pol Pérez Sust   |

- DT 07/03 *Gasto sanitario y envejecimiehto de la población en España* Namkee Ahn, Javier Alonso Meseguer y José A. Herce San Miguel
- DT 01/04 *Métodos de solución de problemas de asignación de recursos sanitarios* Helena Ramalhinho Dias Lourenço y Daniel Serra de la Figuera
- DT 01/05 Licensing of University Inventions: The Role of a Technology Transfer Office Inés Macho-Stadler, David Pérez-Castrillo, Reinhilde Veugelers



### Fundación **BBVA**

Gran Vía, 12 48001 Bilbao Tel.: 94 487 52 52 Fax: 94 424 46 21

Paseo de Recoletos, 10 28001 Madrid Tel.: 91 374 54 00 Fax: 91 374 85 22 informacion@fbbva.es

www.fbbva.es

